

Transmuted Alpha Power Inverse Rayleigh Distribution: Properties and Application

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Abstract

This paper proposes a new three parameter-distribution through the technique known as Transmutation. The proposed distribution is named Transmuted Alpha power inverse Rayleigh Distribution. Several important properties of the distribution are derived. The parameter estimation is also carried out. Two real life data set are used at the end to describe the potential application of proposed model.

Keywords: Transmuted alpha power inverse Rayleigh distribution; Reliability; Entropy; Order statistics.

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1. Introduction

In recent years, there has been an increased interest among researchers to construct new flexible distribution by adding an additional parameter to the baseline distribution. An interesting technique to accomplish the task of parameter addition is transmutation. The Quadratic Rank Transmutation Map (QRTM) was suggested by Shaw and Buckley [1]. They used QRTM to construct non-Gaussian distribution. A number of authors have considered this generalization technique and successfully achieved efficiency over the base distribution. For example “Transmuted Lomax distribution” by Ashour and Eltehiwy [2], “Transmuted Weibull Distribution: A generalization of Weibull Distribution” by Aryal and Tsokos [3], “Transmuted power function distribution” by Haq *et al.* [4], “Transmuted New Weibull Pareto Distribution and its application” by Tahir *et al.* [5], “Transmuted Kumaraswamy Quasi Lindley Distribution with applications” by Elgarhy *et al.* [6], “Transmuted Pareto Distribution” by Merovci and Puka [7], “Transmuted Lindley Geometric distribution and its application” by Merovci and Elbatal [8], “Transmuted Weibull Power Function Distribution: its properties and applications ” by Haq *et al.* [9], “A New Weibull Rayleigh Distribution with application to real life data” by Malik and

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Ahmad [10] etc. In this paper, Alpha Power Inverse Rayleigh Distribution (APIRD) suggested by Malik and Ahmad [11] is generalized using the QRTM. The proposed model is named as Transmuted Alpha Power Inverse Rayleigh Distribution (TAPIRD). Let $f(x)$ and $F(x)$ denote the pdf and cdf of base distribution, the cdf and pdf of generalized distribution are given by Eq. (1) and Eq. (2) respectively.

$$G(x) = (1 + \beta)F(x) - \beta(F(x))^2 \quad ; |\beta| \leq 1. \tag{1}$$

$$g(x) = f(x)[1 + \beta - 2\beta F(x)], \tag{2}$$

where β is the transmuted parameter.

The cdf and pdf of APIRD with scale parameters $\alpha > 0$ and $\lambda > 0$ are given by Eq. (3) and Eq. (4) respectively.

$$F_{APT}(x) = \begin{cases} \frac{(\alpha e^{-\frac{\lambda}{x^2}} - 1)}{(\alpha - 1)} & ; \alpha \neq 1 \\ e^{-\frac{\lambda}{x^2}} & ; \alpha = 1 \end{cases} \tag{3}$$

$$f_{APT}(x) = \begin{cases} \frac{\log \alpha}{(\alpha - 1)} \frac{2\lambda}{x^3} e^{-\frac{\lambda}{x^2}} \alpha e^{-\frac{\lambda}{x^2}} & ; \alpha \neq 1 \\ \frac{2\lambda}{x^2} e^{-\frac{\lambda}{x^2}} & ; \alpha = 1 \end{cases} \tag{4}$$

The main motivation for considering TAPIRD is to add an additional parameter to the base distribution i.e., APIRD so that the flexibility of base distribution can be enhanced. The proposed distribution exhibits more complex shapes of hazard rate function. Therefore, the proposed distributions could be used to model diverse nature of data sets. Also the proposed distribution outperforms some well-known models with respect to two real life data sets. The rest of paper is unfolded as following: in section 2, the pdf and cdf of proposed distribution are introduced, section 3 deals with the reliability analysis of TAPIRD. The expression for mixture representation of the pdf, statistical properties and entropy estimation of TAPIRD are obtained in section 4, 5 and 6 respectively. The expressions for order statistics are derived in section 7. In section 8, the estimation of parameters is accomplished through the technique of maximum likelihood estimation. The simulation study and application of TAPIRD in real life is discussed in section 10 and 11 respectively. Finally some discussions and conclusions are presented at the end.

2. TAPIRD

On substituting Eqs. (3) (4) in Eqs. (1) and (2), we obtain the cdf and pdf of TAPIRD as given by Eqs. (5) and (6) respectively.

$$g(x) = \begin{cases} \frac{\log \alpha}{\alpha - 1} \frac{2\lambda}{x^3} e^{-\lambda/x^2} \alpha^{e^{-\lambda/x^2}} \left[1 + \beta + \frac{2\beta}{\alpha - 1} \left(\frac{2\beta}{\alpha - 1} \alpha^{e^{-\lambda/x^2}} \right) \right]; & \alpha \neq 1 \\ \frac{2\lambda}{x^3} e^{-\lambda/x^2} (1 + \beta - 2\beta e^{-\lambda/x^2}); & \alpha = 1 \end{cases} \quad (5)$$

$$G(x) = \begin{cases} 1 - \frac{\left[(\alpha + \beta) - \alpha (1 - \beta \alpha^{e^{-\lambda/x^2} - 1}) \right] (\alpha - \alpha^{e^{-\lambda/x^2}})}{(\alpha - 1)^2}; & \alpha \neq 1 \\ (1 - \beta) e^{-\lambda/x^2} + 2\beta e^{-\lambda/x^2}; & \alpha = 1 \end{cases} \quad (6)$$

For $\alpha=1$, the Eq. (5) reduces to Transmuted Rayleigh Distribution (TRD) given by [12]. So, for the rest of the paper we will consider only the case for which $\alpha \neq 1$.

The graph of pdf and cdf of TAPIRD for different parameter combination are given by Figs. 1.

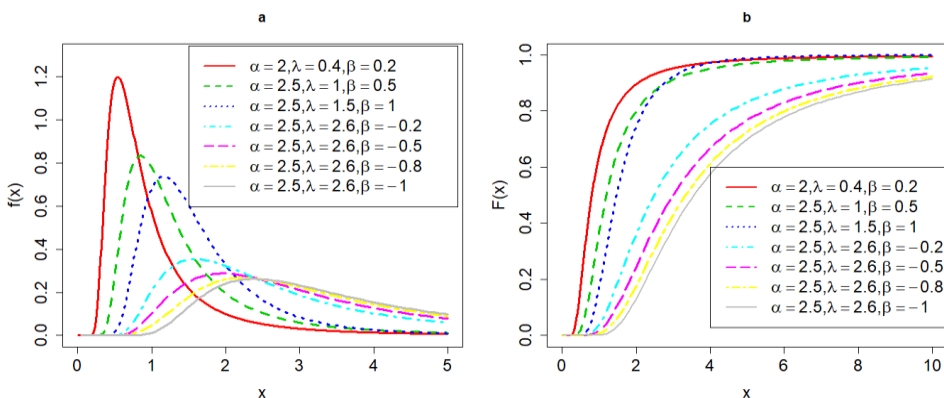


Fig. 1. Plots of (a) pdf and (b) cdf for TAPIRD with several parameter values.

Fig. 1 suggests that the proposed distribution is positively skewed and unimodal. The shape of pdf tends to flatten out for the negative values of the transmuted parameter and becomes peaked for positive values of the transmuted parameter. We further observe from Fig. 2 that the corresponding distribution function is an increasing function of x.

3. Reliability Analysis

3.1. Reliability function

The reliability function of TAPIRD is given as

$$R(x) = \left\{ \frac{\left[(\alpha + \beta) - \alpha(1 - \beta\alpha e^{-\frac{\lambda}{x^2}-1}) \right] (\alpha - \alpha e^{-\frac{\lambda}{x^2}})}{(\alpha - 1)^2} \right\} \tag{7}$$

3.2. Hazard function.

The hazard rate function of TAPIRD is given as

$$h(x) = \frac{2\lambda\beta (\log\alpha) e^{-\frac{\lambda}{x^2}} \alpha e^{-\frac{\lambda}{x^2}}}{x^3 \left\{ (\alpha - 1) - \beta(\alpha e^{-\frac{\lambda}{x^2}} - 1) \right\}} - \frac{2\lambda (\log\alpha) e^{-\frac{\lambda}{x^2}} \alpha e^{-\frac{\lambda}{x^2}}}{x^3 \alpha (1 - \alpha e^{-\frac{\lambda}{x^2}})} \tag{8}$$

3.3. Mean residual life

The mean residual life is defined as

$$\mu(t) = \frac{1}{s(t)} \left[E(t) - \int_0^t xg(x)dx \right] - t \tag{9}$$

Where $E(t) = \sum_{j=0}^{\infty} \frac{\Gamma\left(\frac{1}{2}\right)}{j!(j+1)^{\frac{1}{2}}} \frac{(\log\alpha)^{j+1}}{\alpha - 1} \frac{2\lambda^2}{t^3} e^{-(j+1)\lambda/x^2} \left\{ 1 + \beta - \frac{2\beta(2^j - 1)}{\alpha - 1} \right\}$. (10)

On substituting Eq. (10) and Eq. (5) in Eq. (9) and simplifying, we get

$$\mu(t) = \frac{1}{s(t)} \sum_{j=0}^{\infty} \frac{\lambda^2}{j!(j+1)^{\frac{1}{2}}} \frac{(\log\alpha)^{j+1}}{\alpha - 1} e^{-(j+1)\lambda/x^2} \left\{ 1 + \beta - \frac{2\beta(2^j - 1)}{\alpha - 1} \right\} \gamma\left(\frac{1}{2}, \frac{(j+1)\lambda}{t^2}\right) - t. \tag{11}$$

3.4. Mean waiting time

The mean waiting time is defined as

$$\bar{\mu}(t) = t - \frac{1}{G(t)} \int_0^t xg(x)dx \tag{12}$$

Upon substituting Eq. (5) and Eq. (6) and after simplification yields the below given expression

$$\bar{\mu}(t) = t - \frac{1}{G(t)} \sum_{j=0}^{\infty} \frac{\lambda^2 \Gamma\left(\frac{1}{2}, \frac{(j+1)\lambda}{t^2}\right)}{j!(j+1)^{\frac{1}{2}}} \frac{(\log\alpha)^{j+1}}{\alpha - 1} e^{-(j+1)\lambda/x^2} \left\{ 1 + \beta - \frac{2\beta(2^j - 1)}{\alpha - 1} \right\}. \tag{13}$$

The graph of hazard rate function for different parameter combinations is given by Fig. 2.

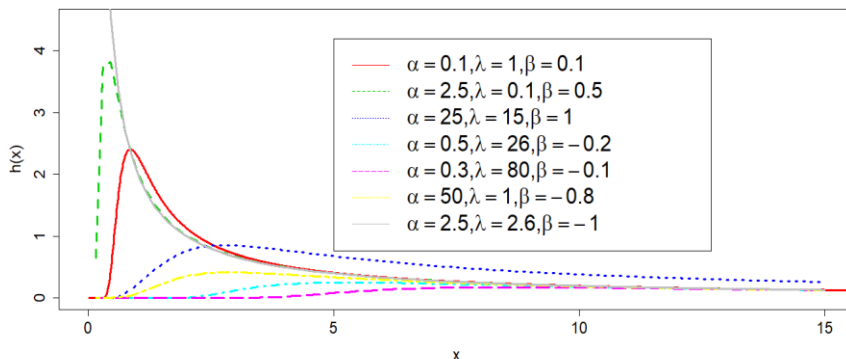


Fig. 2. Plot of hrf of TAPIRD.

Fig. 3 suggests that the proposed distribution is quite flexible in nature and can exhibit variety of shapes such as constant, reverse j-shaped, upside-down bathtub shaped over the parameter space.

4. Mixture Representation

The pdf of TAPIRD can also be expressed as given below

$$g(x) = \sum_{j=0}^{\infty} \frac{(\log \alpha)^{j+1}}{(\alpha - 1) j!} \frac{2\lambda}{x^3} e^{-(j+1)\lambda/x^2} \left\{ 1 + \beta - \frac{2\beta(2^j - 1)}{\alpha - 1} \right\}. \tag{14}$$

The above given mixture representation is very useful in obtaining the properties of TAPIRD.

5. Statistical Properties

In this section, some statistical properties of TAPIRD are discussed.

5.1. Moments

The expression for r^{th} moment about origin of TAPIRD is given as

$$\mu_r' = \sum_{j=0}^{\infty} \frac{\Gamma\left(1 - \frac{r}{2}\right)}{j!(j+1)^{\frac{1-r}{2}}} \frac{(\log \alpha)^{j+1}}{\alpha - 1} \frac{2\lambda^{\frac{r}{2}}}{x^3} e^{-(j+1)\lambda/x^2} \left\{ 1 + \beta - \frac{2\beta(2^j - 1)}{\alpha - 1} \right\}; r < 2 \tag{15}$$

5.2. Incomplete moments about origin

The expression for s^{th} incomplete moment about origin of TAPIRD is given as

$$m_s(k) = \sum_{j=0}^{\infty} \frac{\Gamma\left(1 - \frac{s}{2}, \frac{(j+1)\lambda}{k^2}\right)}{j!(j+1)^{1-\frac{s}{2}}} \frac{(\log \alpha)^{j+1}}{\alpha-1} \frac{2\lambda^2}{x^3} e^{-(j+1)\lambda/x^2} \left\{ 1 + \beta - \frac{2\beta(2^j - 1)}{\alpha - 1} \right\}; r < 2. \quad (16)$$

Theorem 1. The moment generating function and characteristic function of a random variable X following TAPIRD are respectively given by

$$M_X(t) = \sum_{r=0}^{\infty} \sum_{j=0}^{\infty} \frac{t^r \Gamma\left(1 - \frac{r}{2}\right)}{j!r!(j+1)^{1-\frac{r}{2}}} \frac{(\log \alpha)^{j+1}}{\alpha-1} \frac{2\lambda^2}{x^3} e^{-(j+1)\lambda/x^2} \left\{ 1 + \beta - \frac{2\beta(2^j - 1)}{\alpha - 1} \right\}; r < 2. \quad (17)$$

$$\Psi_X(t) = \sum_{r=0}^{\infty} \sum_{j=0}^{\infty} \frac{(it)^r \Gamma\left(1 - \frac{r}{2}\right)}{j!r!(j+1)^{1-\frac{r}{2}}} \frac{(\log \alpha)^{j+1}}{\alpha-1} \frac{2\lambda^2}{x^3} e^{-(j+1)\lambda/x^2} \left\{ 1 + \beta - \frac{2\beta(2^j - 1)}{\alpha - 1} \right\}; r < 2. \quad (18)$$

Proof. The moment generating function is given by

$$M_X(t) = E[e^{tx}]$$

$$M_X(t) = \int_0^{\infty} e^{tx} g(x) dx$$

$$M_X(t) = \int_0^{\infty} \left(1 + tx + \frac{(tx)^2}{2!} + \dots \right) g(x) dx$$

$$M_X(t) = \sum_{l=0}^{\infty} \frac{t^l}{l!} \int_0^{\infty} x^l g(x) dx$$

$$M_X(t) = \sum_{l=0}^{\infty} \frac{t^l}{l!} \mu'_l$$

$$M_X(t) = \sum_{r=0}^{\infty} \sum_{j=0}^{\infty} \frac{t^r \Gamma\left(1 - \frac{r}{2}\right)}{j!r!(j+1)^{1-\frac{r}{2}}} \frac{(\log \alpha)^{j+1}}{\alpha-1} \frac{2\lambda^2}{x^3} e^{-(j+1)\lambda/x^2} \left\{ 1 + \beta - \frac{2\beta(2^j - 1)}{\alpha - 1} \right\}; r < 2.$$

Also, we have

$$\Psi_X(t) = M_X(it)$$

Therefore,

$$\Psi_X(t) = \sum_{r=0}^{\infty} \sum_{j=0}^{\infty} \frac{(it)^r \Gamma\left(1 - \frac{r}{2}\right)}{j!r!(j+1)^{1-\frac{r}{2}}} \frac{(\log \alpha)^{j+1}}{\alpha-1} \frac{2\lambda^2}{x^3} e^{-(j+1)\lambda/x^2} \left\{ 1 + \beta - \frac{2\beta(2^j - 1)}{\alpha - 1} \right\}; r < 2.$$

6. Entropy

The expression Renyi entropy of TAPIRD is given by Eq. (19)

$$RE_X(\delta) = \frac{1}{1-\delta} \sum_{i=0}^{\infty} \sum_{j=0}^{\infty} (-1)^i 2^{\delta-1} \lambda^{\frac{(1-\delta)}{2}} \frac{(\log \alpha)^{j+\delta}}{(\alpha-1)^\delta} \binom{\delta}{i} \left\{ 1 + \beta + \frac{2\beta}{\alpha-1} \right\}^{\delta-i} \left(\frac{2\beta}{\alpha-1} \right)^i (\delta+i)^j \frac{\Gamma\left(\frac{3\delta-1}{2}\right)}{\binom{3\delta-1}{2}^{(\delta+j)}}; \delta \neq 1, \delta > 0. \tag{19}$$

7. Order Statistics

Let X_1, X_2, \dots, X_n be a random sample of size n drawn from TAPIRD and $X_{(1)}, X_{(2)}, \dots, X_{(n)}$ be the corresponding order statistics. Then the pdf of rth order statistics denoted by $g_{r:n}(x)$ is given as

$$g_{r:n}(x) = \frac{n!}{(n-r)!(r-1)!} \left\{ \left(\alpha + \beta - \alpha \left(1 - \beta \alpha^e e^{-\frac{\lambda}{x^2} - 1} \right) \right) \left(\alpha - \alpha^e e^{-\frac{\lambda}{x^2}} \right) \right\}^{n-r} \left\{ (\alpha-1)^2 - \left(\alpha + \beta - \alpha \left(1 - \beta \alpha^e e^{-\frac{\lambda}{x^2} - 1} \right) \right) \right\}^{r-1} \left\{ \left(\alpha - \alpha^e e^{-\frac{\lambda}{x^2}} \right) \right\} \frac{\log \alpha}{(\alpha-1)^{2n}} \frac{2\lambda}{x^3} e^{-\lambda/x^2} \alpha^e e^{-\frac{\lambda}{x^2}} \left\{ 1 + \beta + \frac{2\beta}{\alpha-1} - \left(\frac{2\beta}{\alpha-1} \alpha^e e^{-\frac{\lambda}{x^2}} \right) \right\}. \tag{20}$$

Substituting r = 1 and r = n in Eq. (20) we get the pdf of first and nth order statistics as given by Eqn. (21) and Eqn. (22) respectively.

$$g_{1:n}(x) = n \left\{ \left(\alpha + \beta - \alpha \left(1 - \beta \alpha^e e^{-\frac{\lambda}{x^2} - 1} \right) \right) \left(\alpha - \alpha^e e^{-\frac{\lambda}{x^2}} \right) \right\}^{n-1} \frac{\log \alpha}{(\alpha-1)^{2n}} \frac{2\lambda}{x^3} e^{-\lambda/x^2} \alpha^e e^{-\frac{\lambda}{x^2}} \left\{ 1 + \beta + \frac{2\beta}{\alpha-1} - \left(\frac{2\beta}{\alpha-1} \alpha^e e^{-\frac{\lambda}{x^2}} \right) \right\}. \tag{21}$$

$$g_{n:n}(x) = n \left\{ (\alpha-1)^2 - \left(\alpha + \beta - \alpha \left(1 - \beta \alpha^e e^{-\frac{\lambda}{x^2} - 1} \right) \right) \left(\alpha - \alpha^e e^{-\frac{\lambda}{x^2}} \right) \right\}^{n-1} \frac{\log \alpha}{(\alpha-1)^{2n}} \frac{2\lambda}{x^3} e^{-\lambda/x^2} \alpha^e e^{-\frac{\lambda}{x^2}} \left\{ 1 + \beta + \frac{2\beta}{\alpha-1} - \left(\frac{2\beta}{\alpha-1} \alpha^e e^{-\frac{\lambda}{x^2}} \right) \right\}. \tag{22}$$

8. Parameter Estimation

Let X_1, X_2, \dots, X_n be a random sample of size n drawn from TAPIRD. Then, the log likelihood function is given by Eq. (23)

$$\log l = n \log(\log \alpha) - n \log(\alpha - 1) + n \log(2\lambda) - 3 \sum_{i=1}^n \log x_i - \sum_{i=1}^n \frac{\lambda}{x_i^2} + \sum_{i=1}^n e^{-\lambda/x_i^2} \log \alpha + \log \left\{ 1 + \beta + \frac{2\beta}{\alpha-1} - \left(\frac{2\beta}{\alpha-1} \alpha^e e^{-\frac{\lambda}{x^2}} \right) \right\}. \tag{23}$$

Differentiating Eq. (18) w.r.t. α , λ and β and equating to zero we get the maximum likelihood estimates of the parameters.

9. Simulation Study

In this section, a simulation study has been performed through R software to access the flexibility of proposed model. The values of the parameters are chosen to be $\alpha=3$, $\lambda=3$, and $\beta = 0.3$. The data sets of size 25, 75, 150 and 500 are obtained by using the inverse cdf method and the summary of results is presented in the Table 1 below:

Table 1. Estimates and performance of the distributions for simulated data.

Sample size	Distribution	Estimates			-log l	AIC	BIC	AICC
		α	λ	β				
25	TAPIRD	1.303	4.295	-0.043	50.780	102.561	103.218	102.91
	APIRD	1.397	4.314	-	51.782	107.56	110.000	107.918
	RD	-	4.680	-	51.81	105.636	106.855	105.810
	TRD	-	4.425	0.114	51.79	107.588	110.026	107.762
75	TAPIRD	15.01	2.720	0.622	93.845	183.69	191.427	184.04
	APIRD	2.331	3.340	-	94.325	192.65	196.47	193.00
	RD	-	0.5934	-	94.58	191.172	193.08	191.34
	TRD	-	1.388	0.642	94.21	192.433	196.257	192.607
150	TAPIRD	3.512	3.778	0.551	271.91	541.826	548.85	542.17
	APIRD	0.792	4.315	-	272.38	548.76	554.78	549.11
	RD	-	4.072	-	272.46	546.92	549.93	547.09
	TRD	-	4.352	-0.1354	272.36	548.72	554.74	548.89
500	TAPIRD	0.884	3.968	0.0192	920.01	1837.37	1849.01	1837.72
	APIRD	0.852	3.966	-	920.68	1845.37	1853.80	1845.72
	RD	-	3.814	-	920.82	1843.65	1847.86	1843.82
	TRD	-	3.951	-0.0721	920.69	1845.39	1853.82	1845.57

10. Application

In order to access the flexibility of TAPIRD, two well established models are compared with TAPIRD and criterion such as AIC, BIC and AICC are used to select best model among the compared models for the considered real life data sets. The models used for comparison are Alpha Power Inverse Rayleigh Distribution (APIRD), Rayleigh Distribution (RD) given by Rayleigh [13] and TRD.

Data set 1: The data set 1 represent the survival times (in days) of 72 guinea pigs infected with virulent tubercle bacilli, observed and reported by Bjerkedal [14].

Table 1 shows maximum likelihood estimates and different information measures for models and Fig. 3a shows the plots of the estimated pdfs of APIRD and other competitive models for data set 1.

Table 2. Estimates and performance of the distributions for data set 1.

Distribution	Estimates			-log l	AIC	BIC	AICC
	α	λ	β				
TAPIRD	170.4	0.244	1.00	41.04	88.088	94.51	88.441
APIRD	3918	0.146	-	55.54	115.09	119.37	115.26
RD	-	1.790	-	54.65	111.31	113.45	111.48
TRD	-	0.895	1.00	54.65	113.31	117.59	113.48

Data set 2: The second data set is on the strengths of 1.5 cm glass fibres. The data was originally obtained by workers at the UK National Physical Laboratory and it has been used by Smith and Naylor [15].

Table 2 shows maximum likelihood estimates and different information measures for models and Fig. 3b shows the plots of the estimated pdfs of APIRD and other competitive models for data set 2.

Table 3. Estimates and performance of the distributions for data set 2.

Distribution	Estimates			-log l	AIC	BIC	AICC
	α	λ	B				
TAPIRD	6247	0.0905	0.7877	104.79	215.59	222.42	215.94
APIRD	4557	0.0191	-	109.14	222.28	226.83	222.45
RD	-	0.4629	-	163.75	329.51	331.79	329.69
TRD	-	0.9417	0.3524	140.26	284.53	289.09	284.71

It can be clearly seen from Tables 1 and 2 that TAPIRD has least value of AIC, BIC and AICC for both the data sets. Hence TAPIRD proves to be best model for the given data sets than the compared models.

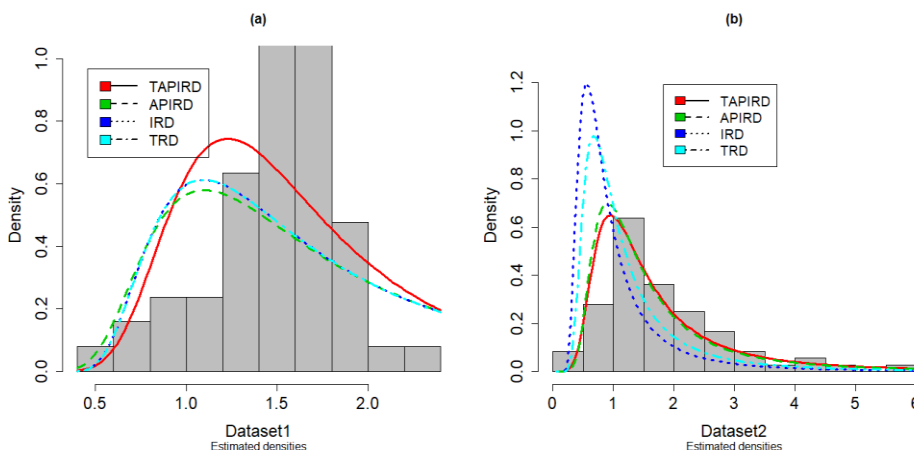


Fig. 3. (a) Plots of the estimated pdf of TAPIRD and other competitive models for data set 1 and (b) plots of the estimated pdf of TAPIRD and other competitive models for data set 2.

11. Conclusion

In this paper, a new model known as Transmuted Alpha Power Inverse Rayleigh Distribution has been successfully defined. The new distribution is more flexible as its hazard rate function exhibits more complex shapes. Two real life data sets were fitted for TAPIRD and compared with three known distributions. The results showed that the TAPIRD is a relatively better model to fit the data than the other three distributions. In simulation study, the same results were obtained. We hope that this distribution attracts wide variety of applications in diverse fields.

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